Bayesian Inference

Summary

- Motivation
- A Posteriori Distribution
- Bayesian Estimation Methods
- Model Selection

Question

Let x be a random variable with values in IR^2 and let y be a linear combination of the x components, corrupted by additive noise:

 $y = x_1 + x_2 + w$

Is it possible to estimate *x* from *y*?



Where is the boat?



Bayes (1702-1761)

Bayesian Inference



- inicial location: p(x) prior
- sensor model: p(y|x)
- final location: p(x|y) a posteriori density function

The final result is a distribution!

A Posteriori Distribution (known model)

How to compute the a posteriori distribution ?



Conjugate Prior

The prior represents the knowledge available about the unknown variables before any observation is made.

It should allow an easy computation of the a posteriori distribution.

A conjugate prior is a prior such that the *a posteriori* distribution has the same analytic expression as the prior, with different values of the parameters.

Exponencial Family

It is easy to obtain conjugate priors if the sensor model p(y|x) belongs to the exponencial family.

Definition: p(y|x) belongs to the **exponential family** if and only if

$$p(y | x) = h(y)g(x)\exp\{t(y)c(x)\}$$
 e $\int p(y | x)dy = 1$

conjugate prior: $p(x) = g(x)^d \exp\{bc(x)\}$

a posteriori density:
$$p(x | y) = g(x)^{\widetilde{d}} \exp\{\widetilde{b}c(x)\}, \quad \widetilde{d} = d + n, \quad \widetilde{b} = b + \sum_{i=1}^{n} t(y_i)$$

Several well known distributions e.g., normal (with known covariance), gamma, binomial, Poisson, belong to the exponential family.

Proof

Let $y = y_1, ..., y_n$ be a sequence of independent observations.

likelihood function

$$p(y \mid x) = g(x)^{n} \prod_{i} h(y_{i}) \exp\{t(y_{i})c(x)\}$$

a posteriori density

$$p(x \mid y) \alpha \quad p(y \mid x)p(x)$$

$$\alpha \quad g(x)^{n} \prod_{i} \quad h(y_{i}) \exp\{t(y_{i})c(x)\} \times g(x)^{d} \exp\{bc(x)\}$$

$$\alpha \quad g(x)^{n+d} \exp\{\left(b + \sum_{i} \quad t(y_{i})\right)c(x)\right\}$$

$$\alpha \quad g(x)^{\widetilde{d}} \exp\{\widetilde{b}c(x)\}$$

Binomial Distribution

The binomial distribution $B(\alpha)$ belongs to the exponential family.

Conjugate prior. $P(\alpha) = c\alpha^{b}(1-\alpha)^{md-b}$ Beta distribution

A posteriori distribution: the same with $\tilde{b} = b + k$, $\tilde{d} = d + 1$

Example: α =.2



Example



This example considers p(x)=N(0,I), p(y/x)=N(x,.04I)

Recursive Computation

Suppose we obtain n independent observations $y=(y_1, ..., y_n)$. Then

$$p(x \mid y) = cp(y_1, y_2 \mid x)p(x) = cp(y_2 \mid x)p(y_1 \mid x)p(x)$$

This suggests the following recursion:

 $p(x | y_{1:k}) \propto p(y_k | x)p(x | y_{1:k-1})$

where $y_{1:k} = (y_1, ..., y_k)$

This procedure is very useful when conjugate priors are used.

A Posteriori Distribution (unknown model)

Let us assume that x depends on an unknown variable θ .

In this case

 $p(x \mid y) = \int p(x \mid \theta) p(\theta \mid y) d\theta$

where

$$p(\theta \mid y) = \frac{p(y \mid \theta)p(\theta)}{p(y)}$$

When the model is unknown the Bayesian approach considers all possible models weighted by their confidence degrees $p(\theta/y)$, instead of using a single (best) model.

MAP and MMSE Estimates

How to obtain an estimate of x from the *a posteriori* distribution ?

MAP estimate (maximum a posteriori)

$$\hat{x} = \underset{x}{\operatorname{argmax}} p(x \mid y) = \underset{x}{\operatorname{argmax}} p(y \mid x) p(x)$$

MMSE estimate (minimum mean squared error)

 $\hat{x} = E\{x \mid y\} = \int x \, p(x \mid y) \, dx$



MAP vs ML



The prior has an important role when there is few data.

(simple rule: the should be 10 observations for each parameter to be estimated.)

Parábolic Fit



Gaussian Variables



Hypothesis: x,y have normal distribution.

Question: what is the distribution of *x* given *y*?

Answer:
$$p(x \mid y) = N(\hat{x}, P)$$

 $\hat{x} = \overline{x} + P_{xy}P_{yy}^{-1}(y - \overline{y})$
 $P = P_{xx} - P_{xy}P_{yy}^{-1}P_{yx}$

Notation: $\overline{a} = E\{a\}, P_{ab} = E\{(a - \overline{a})(b - \overline{b})'\}$

Lemma:

$$\begin{array}{c}
\mathsf{Proof}\\
\begin{bmatrix}A & B\\\\C & D
\end{bmatrix}^{-1} = \begin{bmatrix}E & F\\\\G & H
\end{bmatrix} \qquad E = (A - BD^{-1}C)^{-1}\\F = -EBD^{-1}
\end{array}$$

 $p(x/y) = N(\hat{x}, P)$. The argument of the exponential is

$$q = \begin{bmatrix} \mathbf{x} - \overline{\mathbf{x}} \\ \mathbf{y} - \overline{\mathbf{y}} \end{bmatrix}' \begin{bmatrix} P_{xx} & P_{yx} \\ P_{xy} & P_{yy} \end{bmatrix}^{-1} \begin{bmatrix} \mathbf{x} - \overline{\mathbf{x}} \\ \mathbf{y} - \overline{\mathbf{y}} \end{bmatrix} = (\mathbf{x} - \overline{\mathbf{x}})' \mathbf{E}(\mathbf{x} - \overline{\mathbf{x}}) + 2(\mathbf{x} - \overline{\mathbf{x}})\mathbf{F}(\mathbf{y} - \overline{\mathbf{y}}) + \mathbf{c}$$
$$= \mathbf{x}' \mathbf{E}\mathbf{x} - \overline{\mathbf{x}}) - 2(\mathbf{x} - \overline{\mathbf{x}})(\mathbf{E}\overline{\mathbf{x}} - \mathbf{F}(\mathbf{y} - \overline{\mathbf{y}})) + \mathbf{c}'$$

Comparing with the exponent $\mathcal{O}(\hat{x}, P)$: $\hat{x}'P^{-1}x - 2x'P^{-1}\hat{x} + \hat{x}P^{-1}\hat{x}$

we conclude $P^{-1} = E$, $P^{-1}\hat{x} = E\overline{x} - F(y-\overline{y})$

Therefore,
$$P = (P_{xx} - P_{xy}P_{yy}^{-1}P_{yx})^{-1}, \hat{x} = \bar{x} - P_{xy}P_{yy}^{-1}(y-\bar{y})$$

Example

Let $x \sim N(0,R)$ be a random variable with values in R^2 and y a linear combination of x components, corrupted by white noise:

 $y = x_1 + x_2 + w$

Is it possible to estimate x fom y?



Linear Model

Let us consider a linear model with additive Gaussian noise: y = Cx + v $x \sim N(\overline{x}, \overline{P}), v \sim N(0, Q)$

What is the distribution of *x*, after observing y ?

Answer:
$$p(x | y) = N(\hat{x}, P)$$

$$\hat{x} = \overline{x} + K(y - C\overline{x})$$
 $K = \overline{P}C'S^{-1}$
 $P = (I - KC)\overline{P}$ $S = C\overline{P}C' + R$

This result suggests an incremental update of the parameters when y is a sequence of independent observations.

Bayesian Estimation

Principles:

• The unknown parameters are random variables with known distribution.

• The observations allow to reduce uncertainty of the parameter estimates and to update their distribution. The updated distribution is denoted as **a posteriori distribution**.

• The update is done by the Bayes law.

Notes:

- Bayesian methods provide objective criteria for the design of estimators.
- They have better performance that classic methods when there is few data points.

Difficulties

Inference is more difficult in the following cases:

- invalid data (outliers);
- incomplete data (hidden variables);
- need of model validation/selection;
- *multiple models*

Model Selection



What is the best model?

Model Selection

Let us consider all the available models $M_1, ..., M_c$ to represent a sequence of observations y.

What is the best?

There are several criteria: MV, MAP, MDL, AIC, etc

Occam Razor

In XIV century Occam the following principle:

Choose the simplest model which describes the data with the desired accuracy.

Exercises

1. Let x_1 , ..., x_n be a sequence of independent and identically distributed observations. Knowing that

$$p(x_i \mid \alpha) = \alpha e^{-\alpha x_i}$$
 $p(\alpha) = c e^{-c\alpha}$ $\alpha, x_i > 0$

compute the MAP estimate of a.

2. Consider a signal y_t generated by the model $y_t = a y_{t-1} + b u_t + w_t$ Determine a Bayesian estimate of a,b coefficients assuming that the inputs and outputs $y_1 \dots y_n$, $u_1 \dots u_n$ are available and the noise sequence $w_1 \dots w_n$ consists of uncorrelated variables $w_i \sim N(0,\sigma^2)$.

3. Show that a density p(x)=Cexp[-0.5 (x'Ax+b'x)] is normal $N(\mu,P)$ with $P=A^{-1} e \mu=-0.5 A^{-1} b$.

4. Show that the product of two normal densities $N(\mu_i, P_i)$, i=1,2, is a normal density (apart from a scale factor).

Work

Consider data generated by two probabilistic models

a)
$$\mathbf{x} \sim \mathbf{N}(\mu, \sigma^2)$$
 with known σ^2 and $\mu \sim \mathbf{N}(\mu_0, \sigma_0^2)$
b) $p(x/\alpha) = \alpha e^{-\alpha x}$ $x > 0$, $p(\alpha) = \alpha_0 e^{-\alpha_0 \alpha}$ $\alpha > 0$,

Given an observation x, determine a criteria for the selection of the model.

Characterize the performance of the previous method computing the error probability experimentally.

Bibliography

J. Marques, Reconhecimento de Padrões. Métodos Estatísticos e Neuronais, IST Press, 1999