

Discrete and continuous univariate distributions

$X$ (r.v.)	Values	$P(X = x)$ or $f_X(x)$	$E(X)$	$V(X)$	$M_X(t)$ or $E(X^k)$	$P_X(s)$
Binomial( $n, p$ )	$\{0, \dots, n\}$	$\binom{n}{x} p^x (1-p)^{n-x}$	$np$	$np(1-p)$	$[pe^t + (1-p)]^n$	$(1-p+ps)^n$
HyperG( $N, M, n$ )	$\{\max\{0, n-N+M\}, \dots, \min\{n, M\}\}$	$\frac{\binom{M}{r-x} \binom{N-M}{n-x}}{\binom{N}{n}}$	$n \frac{M}{N}$	$n \frac{M}{N} \frac{N-M}{N-1} \frac{N-n}{N-1}$	not interesting	not interesting
Geometric( $p$ )	$\mathbb{N}$	$(1-p)^{x-1} p$	$\frac{1}{p}$	$\frac{1-p}{p^2}$	$\frac{pe^t}{1-(1-p)e^t}$	$\frac{ps}{1-(1-p)s}$
Geometric $^*$ ( $p$ )	$\mathbb{N}_0$	$(1-p)^x p$	$\frac{1-p}{p}$	$\frac{1-p}{p^2}$	$\frac{pe^t}{1-(1-p)e^t}$	$\frac{ps}{1-(1-p)s}$
NegativeBin( $r, p$ )	$\{r, r+1, \dots\}$	$\binom{x-1}{r-1} p^r (1-p)^{x-r}$	$\frac{r}{p}$	$\frac{r(1-p)}{p^2}$	$\left[ \frac{pe^t}{1-(1-p)e^t} \right]^r$	$\left[ \frac{ps}{1-(1-p)s} \right]^r$
NegativeBin $^*$ ( $r, p$ )	$\{0, 1, \dots\}$	$\binom{y+r-1}{r-1} p^r (1-p)^y$	$\frac{r(1-p)}{p}$	$\frac{r(1-p)}{p^2}$	$\left[ \frac{pe^t}{1-(1-p)e^t} \right]^r$	$\left[ \frac{ps}{1-(1-p)s} \right]^r$
Poisson( $\lambda$ )	$\mathbb{N}_0$	$e^{-\lambda} \frac{\lambda^x}{x!}$	$\lambda$	$\lambda$	$e^{\lambda(e^t-1)}$	$e^{-\lambda(1-s)}$
Uniform ( $\{1, \dots, n\}$ )	$\{1, \dots, n\}$	$\frac{1}{n}$	$\frac{n+1}{2}$	$\frac{n^2-1}{12}$	$\frac{e^t(1-e^{tn})}{n(1-e^t)}$	$\frac{s(1-s^n)}{n(1-s)}$
Beta( $\alpha, \beta$ )	$[0, 1]$	$\frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$	$\frac{\alpha}{\alpha+\beta}$	$\frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$	$1 + \sum_{k=1}^{+\infty} \left( \prod_{r=0}^{k-1} \frac{\alpha+r}{\alpha+\beta+r} \right) \frac{t^k}{k!}$	—
Cauchy( $\mu, \sigma$ )	$\mathbb{R}$	$\frac{1}{\pi\sigma} \frac{1}{1+(\frac{x-\mu}{\sigma})^2}$	nonexistent	nonexistent	nonexistent	—
$\chi_{(n)}^2$	$\mathbb{R}_0^+$	$\frac{(\frac{n}{2})^{\frac{n}{2}}}{\Gamma(\frac{n}{2})} x^{\frac{n}{2}-1} e^{-\frac{x}{2}}$	$n$	$2n$	$\left( \frac{1-t}{2} \right)^{\frac{n}{2}}, t < \frac{1}{2}$	—
Exponential( $\lambda$ )	$\mathbb{R}_0^+$	$\lambda e^{-\lambda x}$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$	$\frac{\lambda}{\lambda-t}, t < \lambda$	—
Gamma( $\alpha, \lambda$ )	$\mathbb{R}_0^+$	$\frac{\lambda^\alpha}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x}$	$\frac{\alpha}{\lambda}$	$\frac{\alpha}{\lambda^2}$	$\left( \frac{\lambda}{\lambda-t} \right)^\alpha, t < \lambda$	—
LogNormal( $\mu, \sigma^2$ )	$\mathbb{R}^+$	$\frac{1}{\sqrt{2\pi\sigma x}} e^{-\frac{(\log \frac{x-\mu}{\sigma})^2}{2\sigma^2}}$	$e^{\mu+\frac{\sigma^2}{2}}$	$(e^{\sigma^2}-1) e^{2\mu+\sigma^2}$	$E(X^k) = e^{k\mu+\frac{k^2\sigma^2}{2}}$	—
Normal( $\mu, \sigma^2$ )	$\mathbb{R}$	$\frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$	$\mu$	$\sigma^2$	$e^{\mu t + \frac{(t\sigma)^2}{2}}$	—
Rayleigh( $\sigma$ )	$\mathbb{R}_0^+$	$\frac{x}{\sigma^2} e^{-\frac{x^2}{2\sigma^2}}$	$\sigma\sqrt{\frac{\pi}{2}}$	$\frac{4-\pi}{2} \sigma^2$	$E(X^k) = (\sqrt{2}\sigma)^k \Gamma(1+\frac{k}{2})$	—
Uniform( $a, \beta$ )	$[a, b]$	$\frac{1}{b-a}$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	$\frac{e^{bt}-e^{at}}{(b-a)t}, t \neq 0$	—
Weibull( $\alpha, \beta$ )	$\mathbb{R}_0^+$	$\frac{\beta}{\alpha} \left(\frac{x}{\alpha}\right)^{\beta-1} e^{-\left(\frac{x}{\alpha}\right)^\beta}$	$\alpha\Gamma\left(1+\frac{1}{\beta}\right)$	$\alpha^2 \left[ \Gamma\left(1+\frac{2}{\beta}\right) - \Gamma\left(1+\frac{1}{\beta}\right)^2 \right]$	$E(X^k) = \alpha^k \Gamma\left(1+\frac{k}{\beta}\right)$	—

$$\Gamma(\alpha) = \int_0^{+\infty} x^{\alpha-1} e^{-x} dx, \alpha > 0; \quad \Gamma(n) = (n-1)!, n \in \mathbb{N}; \quad \Gamma(\alpha+1) = \alpha\Gamma(\alpha), \alpha > 0; \quad \Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}$$

## Relating c.d.f.

$$F_{NegativeBin(r,p)}(x) = 1 - F_{Binomial(x,p)}(r-1)$$

$$F_{Erlang(n,\lambda)}(x) = 1 - F_{Poisson(\lambda x)}(n-1)$$

$$F_{Gamma(\alpha,\beta)}(x) = F_{\chi^2_{(2\alpha)}}(2\beta x)$$

$$F_{Beta(\alpha,\beta)}(x) = 1 - F_{Binomial(\alpha+\beta-1,x)}(\alpha-1)$$

## Moment/probability generating function; moments

$$M_X(t) = E(e^{tX})$$

$$E(X^k) = \left. \frac{d^k M_X(t)}{dt^k} \right|_{t=0}$$

$$P_X(s) = E(s^X); \quad P(X=k) = \left. \frac{1}{k!} \times \frac{d^k P_X(s)}{ds^k} \right|_{s=0}$$

$$E[X(X-1)\cdots(X-k+1)] = \left. \frac{d^k P_X(s)}{ds^k} \right|_{s=1}, \quad k \in \mathbb{N}$$

$$E(X) = \int_0^{+\infty} [1 - F_X(x)] dx, \quad \text{for } X \geq 0$$

$$E(X^k) = \int_0^{+\infty} kx^{k-1} [1 - F_X(x)] dx, \quad \text{for } X \geq 0$$

$$SC(X) = \frac{E\{[X-E(X)]^3\}}{[SD(X)]^3}$$

$$KC(X) = \frac{E\{[X-E(X)]^4\}}{[SD(X)]^4} - 3$$

## Multinomial distribution

$$P(N_1 = n_1, \dots, N_d = n_d) = \frac{n!}{\prod_{i=1}^d n_i!} \times \prod_{i=1}^d p_i^{n_i}$$

$$\{(n_1, \dots, n_d) \in \mathbb{N}_0^d : \sum_{i=1}^d n_i = n\}$$

$$M_{N_1, \dots, N_{d-1}}(t_1, \dots, t_{d-1}) = \left[ \left( \sum_{i=1}^{d-1} p_i e^{t_i} \right) + p_d \right]^n$$

$$N_i \sim \text{Binomial}(n, p_i); \quad \text{Cov}(N_i, N_j) = -n p_i p_j, \quad i \neq j$$

$$M_{\underline{X}}(\underline{t}) = E[\exp(\sum_{i=1}^n t_i X_i)]$$

$$E\left(\prod_{i=1}^n X_i^{k_i}\right) = \left. \frac{\partial^{\sum_{i=1}^n k_i} M_{\underline{X}}(\underline{t})}{\partial t_1^{k_1} \dots \partial t_n^{k_n}} \right|_{\underline{t}=0}$$

## Functions of r.v.

$$F_{Y=g(X)}(y) = P[X \in g^{-1}((-\infty, y])]$$

$$f_{Y=g(X)}(y) = f_X[g^{-1}(y)] \times \left| \frac{dg^{-1}(y)}{dy} \right|$$

## Hierarchical models resulting from mixtures

$$P(X=x) = \sum_y P(X=x|Y=y) \times P(Y=y)$$

$$P(X=x) = \int_{\mathbb{R}_Y} P(X=x|Y=y) \times f_Y(y) dy$$

$$E[g(X)] = E\{E[g(X)|Y]\}$$

$$V[g(X)] = V\{E[g(X)|Y]\} + E\{V[g(X)|Y]\}$$

## Functions of random vectors

$$F_{\underline{Y}=g(\underline{X})}(\underline{y}) = P[\underline{X} \in \underline{g}^{-1}(\prod_{i=1}^m (-\infty, y_i])]$$

$$f_{\underline{Y}=g(\underline{X})}(\underline{y}) = f_{\underline{X}}[\underline{g}^{-1}(\underline{y})] \times |J(\underline{y})|$$

$$J(\underline{y}) = \begin{vmatrix} \frac{\partial g_1^{-1}(\underline{y})}{\partial y_1} & \dots & \frac{\partial g_1^{-1}(\underline{y})}{\partial y_n} \\ \vdots & \dots & \vdots \\ \frac{\partial g_n^{-1}(\underline{y})}{\partial y_1} & \dots & \frac{\partial g_n^{-1}(\underline{y})}{\partial y_n} \end{vmatrix}$$

$$f_{X+Y}(z) = \int_{-\infty}^{+\infty} f_{X,Y}(z-y, y) dy$$

$$f_{X-Y}(u) = \int_{-\infty}^{+\infty} f_{X,Y}(u+y, y) dy$$

$$f_{XY}(v) = \int_{-\infty}^{+\infty} f_{X,Y}(v/y, y) \times \frac{1}{|y|} dy$$

$$f_{X/Y}(w) = \int_{-\infty}^{+\infty} f_{X,Y}(wy, y) \times |y| dy$$

## Order statistics

$$P[X_{(n-k+1)} > x] = 1 - F_{Binomial(n, 1-F_X(x))}(k-1)$$

$$f_{X_{(1), \dots, X_{(n)}}}(x_{(1)}, \dots, x_{(n)}) = n! \times \prod_{i=1}^n f_X(x_{(i)})$$

$$F_{X_{(i)}}(x) = 1 - F_{Binomial(n, F_X(x))}(i-1)$$

$$f_{X_{(i)}}(x) = \frac{n!}{(i-1)!(n-i)!} [F_X(x)]^{i-1} [1 - F_X(x)]^{n-i} f_X(x)$$

$$f_{(X_{(i)}, X_{(j)})}(x, y) = \frac{n!}{(i-1)!(j-i-1)!(n-j)!} [F_X(x)]^{i-1} [F_X(y) - F_X(x)]^{j-i-1} [1 - F_X(y)]^{n-j} f_X(x) f_X(y), \quad x < y$$

## Cap. 0

### BERNOULLI PROCESS

$$\{X_n : n \in \mathbb{N}\} \stackrel{i.i.d.}{\sim} \text{Bernoulli}(p)$$

$$S_n = \sum_{i=1}^n X_i, \quad T_k = \min\{n \in \mathbb{N} : S_n = k\}$$

$$\{U_k = T_k - T_{k-1} : k \in \mathbb{N}\} \stackrel{i.i.d.}{\sim} \text{Geometric}(p)$$

$$S_m | S_n = k \sim \text{HyperG}(n, m, k), \quad 0 \leq m \leq n, \quad 0 \leq k \leq n$$

## Cap. 1

### A FEW PROPERTIES OF THE EXPONENTIAL DISTRIBUTION

$$\begin{aligned}
 X_i &\stackrel{\text{indep}}{\sim} \text{Exponential}(\lambda_i), \quad i = 1, \dots, n \Rightarrow & \text{a) } \min_{i=1, \dots, n} \{X_i\} &\sim \text{Exponential}(\sum_{i=1}^n \lambda_i) \\
 \text{b) } P(X_j = \min_{i=1, \dots, n} \{X_i\}) &= \frac{\lambda_j}{\sum_{i=1}^n \lambda_i} & \text{c1) } \sum_{i=1}^n X_i &\sim \text{Hypo-exp.}(\lambda_1, \dots, \lambda_n), \quad \lambda_i \neq \lambda_j \quad (i \neq j) \\
 \text{c2) } f_{\sum_{i=1}^n X_i}(x) &= \sum_{i=1}^n C_{i,n} \times \lambda_i e^{-\lambda_i x}, \quad C_{i,n} = \prod_{j \neq i} \frac{\lambda_j}{\lambda_j - \lambda_i} \\
 \text{d1) } X &\sim \text{Hyper-exp.}(\lambda_1, \dots, \lambda_n; p_1, \dots, p_n), \quad \lambda_i \neq \lambda_j \quad (i \neq j) & \text{d2) } f_X(x) &= \sum_{i=1}^n p_i f_{X_i}(x)
 \end{aligned}$$

### HOMOGENOUS PP (independent and stationary increments)

$$\begin{aligned}
 \{N(t) : t \geq 0\} &\sim PP(\lambda) & S_n &= \min\{t \geq 0 : N(t) = n\}; \quad N(t) \geq n \Leftrightarrow S_n \leq t \\
 N(t) &\sim \text{Poisson}(\lambda t) & S_n &\sim \text{Erlang}(n, \lambda); \quad F_{S_n}(t) = 1 - F_{\text{Poisson}(\lambda t)}(n-1) \\
 (N(s) | N(t) = n) &\sim \text{Binomial}(n, s/t), \quad 0 < s < t \\
 \{N_i(t) : t \geq 0\} &\stackrel{\text{indep}}{\sim} PP(\lambda_i), \quad i = 1, 2 \Rightarrow & P[S_n^{(1)} < S_m^{(2)}] &= 1 - F_{\text{Binomial}(n+m-1, \frac{\lambda_1}{\lambda_1+\lambda_2})}(n-1) \\
 \{X_n = S_n - S_{n-1} : n \in \mathbb{N}\} &\stackrel{\text{i.i.d.}}{\sim} \text{Exp}(\lambda) & (S_1 | N(t) = 1) &\sim \text{Uniform}(0, t) \\
 Y_i &\stackrel{\text{i.i.d.}}{\sim} \text{Uniform}(0, t), \quad i = 1, \dots, n & (S_1, \dots, S_n | N(t) = n) &\sim (Y_{(1)}, \dots, Y_{(n)})
 \end{aligned}$$

$N_1(t) = \#$ registered events in  $(0, t]$ , under the non-homogeneous Bernoulli splitting mechanism associated with a  $p : \mathbb{R}_0^+ \rightarrow [0, 1]$

$$N_1(t) \sim \text{Poisson}\left(\lambda \int_0^t p(s) ds\right)$$

### NON-HOMOGENOUS PP (independent increments)

$$\begin{aligned}
 \{N(t) : t \geq 0\} &\sim NHPP(\lambda(t)) & m(t) &= \int_0^t \lambda(z) dz \\
 N(t+s) - N(s) &\sim \text{Poisson}(m(t+s) - m(s)) & (N(s) | N(t) = n) &\sim \text{Binomial}(n, \frac{m(s)}{m(t)}), \quad 0 < s < t, \quad n \in \mathbb{N} \\
 f_{S_n}(t) &= \lambda(t) e^{-m(t)} \frac{[m(t)]^{n-1}}{(n-1)!}, \quad n \in \mathbb{N} & P(X_{n+1} = S_{n+1} - S_n > t) &= \int_0^{+\infty} \lambda(s) e^{-m(t+s)} \frac{[m(s)]^{n-1}}{(n-1)!} ds, \quad n \in \mathbb{N} \\
 Y_i &\stackrel{\text{i.i.d.}}{\sim} Y, \quad \text{where } P(Y \leq u) = \frac{m(u)}{m(t)}, \quad \text{for } 0 \leq u \leq t & (S_1, \dots, S_n | N(t) = n) &\sim (Y_{(1)}, \dots, Y_{(n)})
 \end{aligned}$$

### CONDITIONAL PP (stationary increments)

$$\{N(t) : t \geq 0\} \sim \text{ConditionalPP}(G) \quad P[N(t+s) - N(s) = n] = \int_0^{+\infty} \frac{e^{-\lambda t} (\lambda t)^n}{n!} dG(\lambda)$$

### COMPOUND PP (independent and stationary increments)

$$\{X(t) = \sum_{i=1}^{N(t)} Y_i : t \geq 0\} \sim \text{CompoundPP}(\lambda, Y) \quad E[X(t)] = \lambda t \times E(Y); \quad V[X(t)] = \lambda t \times E(Y^2)$$

## Cap. 2

### RENEWAL PROCESSES

$$\begin{aligned}
 \{N(t) : t \geq 0\} &\sim RP \text{ with inter-renewal distribution } F \text{ and mean } \mu; & F_n(t) &= P(S_n \leq t) \\
 P[N(t) = n] &= F_n(t) - F_{n+1}(t) & m(t) &= E[N(t)] = \sum_{n=1}^{+\infty} F_n(t) \\
 m(t) &= F(t) + \int_0^t m(t-x) dF(x) & \tilde{m}(s) &= \int_0^{+\infty} e^{-st} dm(t) = \frac{\tilde{F}(s)}{1-\tilde{F}(s)} \\
 H(t) &= D(t) + \int_0^t H(t-x) dF(x) & |D(t)| < +\infty &\Rightarrow H(t) = D(t) + \int_0^t D(t-x) dm(x) \\
 H(t) &= D(t) + \int_0^t D(t-x) dm(x) & \tilde{D}(s) &= \int_0^{+\infty} e^{-st} dD(t) \text{ exists; } \quad \tilde{H}(s) = \frac{\tilde{D}(s)}{1-\tilde{F}(s)}
 \end{aligned}$$

IMPORTANT LAPLACE TRANSFORMS

$f(t)$	$f^*(s) = \int_0^\infty e^{-st} f(t) dt$	$g(t)$	$g^*(s) = \int_0^\infty e^{-st} g(t) dt$
1	$1/s$	$af(t) + bh(t)$	$af^*(s) + bh^*(s)$
$t^n$	$n!/s^{n+1}$	$\frac{df(t)}{dt}$	$sf^*(s) - f(0)$
$\frac{t^{n-1}e^{-at}}{(n-1)!}$	$1/(s+a)^n$	$e^{-at} f(t)$	$f^*(s+a)$
$\sin(at)$	$a/(s^2+a^2)$	$\int_0^t f(u) du$	$f^*(s)/s$
$\frac{e^{-at}-e^{-bt}}{b-a}$	$1/[(s+a)(s+b)]$	$t^n f(t)$	$(-1)^n \frac{d^n}{ds^n} f^*(s)$

LIMIT THEOREMS ET AL.

$$P[N(t) < n] \simeq \Phi\left(\frac{n-t/\mu}{\sqrt{t\sigma^2/\mu^3}}\right)$$

SLLN for renewal processes:  $\frac{N(t)}{t} \xrightarrow{w.p.1} \frac{1}{\mu}$

Elementary renewal theorem:  $\lim_{t \rightarrow +\infty} \frac{m(t)}{t} = \frac{1}{\mu}$

Key renewal th.:  $D(t)$  dRi,  $F$  not lattice,  $H(t) = D(t) + \int_0^t D(t-x) dm(x) \Rightarrow \lim_{t \rightarrow +\infty} H(t) = \frac{1}{\mu} \int_0^{+\infty} D(y) dy$

Blackwell's th.:  $F$  not lattice  $\Rightarrow \lim_{t \rightarrow +\infty} [m(t+a) - m(t)] = \frac{a}{\mu}$

$F$  lattice with period  $d \Rightarrow \lim_{n \rightarrow +\infty} E[\text{number of renewals at } nd] = \frac{d}{\mu}$

RECURRENCE TIMES

$$A(t) = t - S_{N(t)}, \quad Y(t) = S_{N(t)+1} - t, \quad X_{N(t)+1} = A(t) + Y(t)$$

$$E[S_{N(t)+1}] = \mu \times [m(t) + 1]$$

$$X_{N(t)+1} \geq_{st} X_i, \quad i \in \mathbb{N} \text{ (inspection paradox)}$$

$$\frac{A(t)}{t} \xrightarrow{w.p.1} 0, \quad \lim_{t \rightarrow 0} \frac{E[Y(t)]}{t} = 0$$

$$\lim_{t \rightarrow +\infty} E[Y(t)] = \frac{E(X^2)}{2\mu}$$

$$\lim_{t \rightarrow +\infty} E[A(t)] = \frac{E(X^2)}{2\mu}$$

$$\lim_{t \rightarrow +\infty} E[X_{N(t)+1}] = \frac{E(X^2)}{\mu} \geq E(X)$$

$$\lim_{t \rightarrow +\infty} P[Y(t) \leq x] = \lim_{t \rightarrow +\infty} P[A(t) \leq x] = F_e(x) = \frac{\int_0^x [1-F(u)] du}{\mu} \text{ (equilibrium distribution)}$$

REWARD RENEWAL PROCESSES

$$\{(X_n, R_n) : n \in \mathbb{N}\} \stackrel{i.i.d.}{\sim} (X, R), \quad R(t) = \sum_{n=1}^{N(t)} R_n$$

$$E(X), E(R) < +\infty \Rightarrow$$

a)  $\frac{R(t)}{t} \xrightarrow{w.p.1} \frac{E(R)}{E(X)}$

b)  $\lim_{t \rightarrow +\infty} \frac{E[R(t)]}{t} = \frac{E(R)}{E(X)}$

ALTERNATING RENEWAL PROCESSES

$$\{(U_n, D_n) : n \in \mathbb{N}\} \stackrel{i.i.d.}{\sim} (U, D), \quad Z(t) = \begin{cases} 1, & \text{if } \exists n \in \mathbb{N} : S_n = \sum_{i=1}^n X_i = \sum_{i=1}^n (U_i + D_i) \leq t < S_n + U_{n+1} \\ 0, & \text{if } \exists n \in \mathbb{N} : S_n + U_{n+1} \leq t < S_{n+1} \end{cases}$$

$$E(U_n + D_n) < +\infty, F \text{ not lattice} \Rightarrow \lim_{t \rightarrow +\infty} P[Z(t) = 1] = \frac{E(U)}{E(U)+E(D)}$$

DELAYED RENEWAL PROCESSES

$$\{N_D(t) : t \geq 0\}, \text{ with } X_1 \sim G, X_i \sim F, i = 2, 3, \dots, \text{ and } (G \star F_{n-1})(t) = P(S_n \leq t) = \int_0^t G(t-x) dF_{n-1}(x)$$

$$P[N_D(t) = n] = P(S_n \leq t) - P(S_{n+1} \leq t) = (G \star F_{n-1})(t) - (G \star F_n)(t)$$

$$m_D(t) = \sum_{n=1}^{+\infty} (G \star F_{n-1})(t)$$

$$\tilde{m}_D(s) = \frac{\tilde{G}(s)}{1-F(s)}$$

REGENERATIVE PROCESSES

$$\{X(t) : t \geq 0\}, \text{ with state space } \mathbb{N}_0 \text{ and } S_1 \sim F$$

$$U_j = \text{time spent in state } j \text{ during } [0, S_1]$$

$$F \text{ not lattice, } E(S_1) < +\infty \Rightarrow \lim_{t \rightarrow +\infty} P[X(t) = j] = \frac{E(U_j)}{E(S_1)} = P_j$$

$$E(S_1) < +\infty \Rightarrow \lim_{t \rightarrow +\infty} \frac{\text{amount of time state } j \text{ during } (0,t)}{t} \stackrel{w.p.1}{=} P_j$$

### Cap. 3

#### DEFINITIONS AND EXAMPLES

$$\{X_n : n \in \mathbb{N}_0\}, \quad P(X_{n+1} = j \mid X_n = i, X_{n-1} = i_{n-1}, \dots, X_0 = i_0) = P(X_{n+1} = j \mid X_n = i) = P(X_1 = j \mid X_0 = i) = P_{ij}$$

$$\mathbf{P} = [P_{ij}]_{i,j \in \mathcal{S}}, \quad \underline{\alpha} = [\alpha_i]_{i \in \mathcal{S}} = [P(X_0 = i)]_{i \in \mathcal{S}}$$

#### CHAPMAN-KOLMOGOROV EQUATIONS; MARGINAL AND JOINT DISTRIBUTIONS

$$P_{ij}^{n+m} = P(X_{n+m} = j \mid X_0 = i) = \sum_{k \in \mathcal{S}} P_{ik}^n P_{kj}^m \quad \mathbf{P}^n = [P_{ij}^n]_{i,j \in \mathcal{S}}, \quad \underline{\alpha}^n = [P(X_n = j)]_{j \in \mathcal{S}} = \underline{\alpha} \mathbf{P}^n$$

$$P(X_{n_1} = i_{n_1}, \dots, X_{n_k} = i_{n_k}) = \left( \sum_{i \in \mathcal{S}} \alpha_i \times P_{i, i_{n_1}}^{n_1} \right) \times \prod_{j=2}^k P_{i_{n_{j-1}}, i_{n_j}}^{n_j - n_{j-1}}, \quad 0 \leq n_1 < n_2 < \dots < n_k, \quad i_{n_1}, \dots, i_{n_k} \in \mathcal{S}$$

#### CLASSIFICATION OF STATES; RECURRENT AND TRANSIENT STATES

$$f_{ij}^n = P(X_n = j, X_{n-1} \neq j, \dots, X_1 \neq j \mid X_0 = i) \quad T_i = \min\{n \in \mathbb{N} : X_n = i \mid X_0 = i\}$$

$$f_{ij} = \sum_{n=1}^{+\infty} f_{ij}^n, \quad f_i \equiv f_{ii} = \sum_{n=1}^{+\infty} f_{ii}^n = P(T_i < +\infty) \quad \text{recurrent if } f_i = 1, \quad \text{transient if } f_i < 1$$

$$\mu_{ii} = E(T_i) = \sum_{n=1}^{+\infty} n \times f_{ii}^n \quad (i \text{ recurrent}) \quad \text{positive rec. if } \mu_{ii} < +\infty, \quad \text{null rec. if } \mu_{ii} = +\infty$$

#### LIMIT BEHAVIOR OF IRREDUCIBLE (APERIODIC) MARKOV CHAINS

$$P_j = \sum_{i \in \mathcal{S}} P_i P_{ij}, \quad j \in \mathcal{S} \quad (\text{stationary dist.}); \quad \lim_{n \rightarrow +\infty} P_{ij}^n = \pi_j > 0, \quad i, j \in \mathcal{S} \quad (\text{if all states are posit. rec. aper.})$$

$$\pi_j = \sum_{i \in \mathcal{S}} \pi_i P_{ij}, \quad j \in \mathcal{S}, \quad \sum_{j \in \mathcal{S}} \pi_j = 1; \quad \mu_{jj} = \frac{1}{\pi_j}; \quad \underline{\pi} = \underline{\pi} \mathbf{P}, \quad \underline{\pi} \underline{1} = 1; \quad \underline{\pi} = [\pi_j]_{j \in \mathcal{S}} = \underline{1} \times (\mathbf{I} - \mathbf{P} + \mathbf{ONE})^{-1}$$

#### MARKOV CHAINS WITH COSTS/REWARDS

$$\lim_{N \rightarrow +\infty} \frac{1}{N+1} E \left[ \sum_{n=0}^N c(X_n) \mid X_0 = i \right] = \sum_{j \in \mathcal{S}} \pi_j c(j) = \underline{\pi} \times \underline{c} \quad (\text{long-run expected cost per time unit})$$

$$\phi(i) = E \left[ \sum_{n=0}^{+\infty} \alpha^n c(X_n) \mid X_0 = i \right] \quad (\text{expected total discounted cost incurred over... , starting at state } i)$$

$$\phi(i) = c(i) + \alpha \sum_{j \in \mathcal{S}} P_{ij} \phi(j), \quad i \in \mathcal{S} \quad \underline{\phi} = [\phi(i)]_{i \in \mathcal{S}} = (\mathbf{I} - \alpha \mathbf{P})^{-1} \times \underline{c}$$

#### TIME REVERSIBLE MARKOV CHAINS

$$\{X_{n-m} : m \in \mathbb{Z}\}, \quad Q_{ij} = \frac{\pi_j \times P_{ji}}{\pi_i} \quad \pi_i \times P_{ij} = \pi_j \times P_{ji}, \quad i, j \in \mathcal{S}$$

$$P_{i, i_1} \times P_{i_1, i_2} \times \dots \times P_{i_k, i} = P_{i, i_k} \times \dots \times P_{i_2, i_1} \times P_{i_1, i}, \quad \text{for any } i, i_1, i_2, \dots, i_k, k \in \mathbb{N} \quad (\text{Kolmogorov's criterion})$$

#### BRANCHING PROCESSES

$$\{X_n : n \in \mathbb{N}_0\}, \quad X_0 = 1, \quad X_n = \sum_{l=1}^{X_{n-1}} Z_l, \quad n \in \mathbb{N}, \quad \{Z_l : l \in \mathbb{N}\} \text{ i.i.d., non-negat. r.v., } P_j = P(Z_l = j), \quad j \in \mathbb{N}_0$$

$$P_n(s) = E(s^{X_n}), \quad P(s) = E(s^{Z_l}), \quad s \in [0, 1] \quad P_{n+1}(s) = P_n[P(s)] = P[P_n(s)], \quad n \in \mathbb{N}, \quad s \in [0, 1]$$

$$\mu = E(Z_l), \quad \sigma^2 = V(Z_l), \quad E(X_n \mid X_0 = 1) = \mu^n \quad V(X_n \mid X_0 = 1) = \begin{cases} \sigma^2 \mu^{n-1} \times \frac{\mu^n - 1}{\mu - 1}, & \text{if } \mu \neq 1 \\ n\sigma^2, & \text{if } \mu = 1, \end{cases}$$

$$\pi = \lim_{n \rightarrow +\infty} P(X_n = 0 \mid X_0 = 1) \quad \text{if } \mu \leq 1, \pi = 1; \quad \text{if } \mu > 1, \pi = \sum_{j=0}^{+\infty} \pi^j \times P_j$$

#### FIRST PASSAGE TIMES; ABSORPTION PROBABILITIES

$$f_{ij}^n = P(X_n = j, X_{n-1} \neq j, \dots, X_1 \neq j \mid X_0 = i) \quad \underline{f}_j^n = [f_{ij}^n]_{i \in \mathcal{S}}$$

$$f_{ij}^n = \begin{cases} P_{ij}, & n = 1 \\ \sum_{k \neq j} P_{ik} f_{kj}^{n-1}, & n = 2, 3, \dots \end{cases} \quad \underline{f}_j^n = \begin{cases} \underline{f}_j^1 = [P_{ij}]_{i \in \mathcal{S}}, & n = 1 \\ {}^{(j)}\mathbf{P} \times \underline{f}_j^{n-1} = [{}^{(j)}\mathbf{P}]^{n-1} \times \underline{f}_j^1, & n = 2, 3, \dots \end{cases}$$

$$\mathcal{S} = T \cup C_1 \cup C_2 \cup \dots, \quad \mathbf{Q} = [Q_{ij}]_{i,j \in T}, \quad \mathbf{R} = [P_{kl}]_{k \in T, l \in \bar{T}}, \quad \tau = \inf\{n \in \mathbb{N}_0 : X_n \notin T\}, \quad u_{ik} = P(X_\tau = k \mid X_0 = i)$$

$$\mathbf{U} = [u_{ik}]_{i \in T, k \in \bar{T}} = (\mathbf{I} - \mathbf{Q})^{-1} \times \mathbf{R} \quad [E(\tau \mid X_0 = i)]_{i \in T} = (\mathbf{I} - \mathbf{Q})^{-1} \times \underline{1}$$

$$\left[ E\left[ \sum_{n=0}^{\tau-1} g(X_n) \mid X_0 = i \right] \right]_{i \in T} = (\mathbf{I} - \mathbf{Q})^{-1} \times \underline{g}$$

## Cap. 4

### DEFINITIONS AND EXAMPLES

$$\{X(t) : t \geq 0\}, \quad P[X(t+s) = j \mid X(s) = i, X(u) = x(u), 0 \leq u < s] = P[X(t+s) = j \mid X(s) = i]$$

$$P[X(t+s) = j \mid X(s) = i] = P[X(t) = j \mid X(0) = i] = P_{ij}(t), \quad \mathbf{P}(t) = [P_{ij}(t)]_{i,j \in \mathcal{S}}, \quad \underline{\alpha} = [\alpha_i]_{i \in \mathcal{S}} = [P[X(0) = i]]_{i \in \mathcal{S}}$$

### PROPERTIES OF THE TRANSITION MATRIX; CHAPMAN-KOLMOGOROV EQUATIONS

$$P_{ij}(t+s) = \sum_{k \in \mathcal{S}} P_{ik}(t) \times P_{kj}(s) \quad \mathbf{P}(t+s) = \mathbf{P}(t) \times \mathbf{P}(s) = \mathbf{P}(s) \times \mathbf{P}(t)$$

$$P[X(t) = j] = \sum_{i \in \mathcal{S}} \alpha_i \times P_{ij}(t), \quad j \in \mathcal{S} \quad [P[X(t) = j]]_{j \in \mathcal{S}} = \underline{\alpha} \times \mathbf{P}(t)$$

$$q_{ij} = \nu_i \times P_{ij}, \quad i \neq j \quad \lim_{h \rightarrow 0^+} \frac{P_{ij}(h)}{h} = q_{ij}, \quad i \neq j; \quad \lim_{h \rightarrow 0^+} \frac{1 - P_{ii}(h)}{h} = \nu_i$$

$$\frac{dP_{ij}(t)}{dt} = \sum_{k \neq i} q_{ik} P_{kj}(t) - \nu_i P_{ij}(t) \quad (\text{backward eq.}) \quad \frac{dP_{ij}(t)}{dt} = \sum_{k \neq j} P_{ik}(t) q_{kj} - P_{ij}(t) \nu_j \quad (\text{forward eq.})$$

$$r_{ij} = \begin{cases} q_{ij}, & i \neq j \\ -\nu_i, & i = j, \end{cases} \quad \mathbf{R} = [r_{ij}]_{i,j \in \mathcal{S}} \quad \frac{d\mathbf{P}(t)}{dt} = \left[ \frac{dP_{ij}(t)}{dt} \right]_{i,j \in \mathcal{S}} = \mathbf{R} \times \mathbf{P}(t) = \mathbf{P}(t) \times \mathbf{R}$$

### COMPUTING THE TRANSITION MATRIX: FINITE STATE SPACE

$$\mathbf{P}(t) = e^{\mathbf{R}t} = \sum_{n=0}^{+\infty} \frac{\mathbf{R}^n t^n}{n!} = \lim_{n \rightarrow +\infty} (\mathbf{I} + \frac{\mathbf{R}t}{n})^n$$

### COMPUTING THE TRANSITION MATRIX: INFINITE STATE SPACE

$$P_{ij}^*(s) = \int_0^{+\infty} e^{-st} P_{ij}(t) dt, \quad i, j \in \mathcal{S}$$

$$\int_0^{+\infty} e^{-st} \frac{dP_{ij}(t)}{dt} dt = s \times P_{ij}^*(s) - P_{ij}(0) = \sum_{k \neq j} P_{ik}^*(s) \times q_{kj} - P_{ij}^*(s) \times \nu_j$$

### BIRTH AND DEATH PROCESSES

$$\frac{dP_{0j}(t)}{dt} = \lambda_0 P_{1j}(t) - \lambda_0 P_{0j}(t), \quad j \in \mathbb{N}_0; \quad \frac{dP_{ij}(t)}{dt} = \lambda_i P_{i+1,j}(t) + \mu_i P_{i-1,j}(t) - (\lambda_i + \mu_i) P_{ij}(t), \quad i \in \mathbb{N}, \quad j \in \mathbb{N}_0 \quad (\text{b. eq.})$$

$$\frac{dP_{i0}(t)}{dt} = P_{i1}(t) \mu_1 - P_{i0}(t) \lambda_0, \quad i \in \mathbb{N}_0; \quad \frac{dP_{ij}(t)}{dt} = P_{i,j-1}(t) \lambda_{j-1} + P_{i,j+1}(t) \mu_{j+1} - P_{ij}(t) (\lambda_j + \mu_j), \quad i \in \mathbb{N}_0, \quad j \in \mathbb{N} \quad (\text{f. eq.})$$

$$P_j(t) \equiv P[X(t) = j \mid X(0) = i] \quad P(z, t) = E[z^{X(t)} \mid X(0) = i], \quad |z| \leq 1$$

$$\sum_{j \in \mathcal{S}} z^j \times \frac{dP_j(t)}{dt} = \frac{\partial P(z, t)}{\partial t} \quad \frac{\partial P(z, t)}{\partial z} = \sum_{j \in \mathcal{S}} j z^{j-1} \times P_j(t) = \sum_{j \in \mathcal{S}} (j+1) z^j \times P_{j+1}(t)$$

$$\sum_{j \in \mathcal{S}} z^j \times \frac{dP_j(t)}{dt} = \sum_{j \in \mathcal{S}} z^j \times [P_{j-1}(t) \lambda_{j-1} + P_{j+1}(t) \mu_{j+1} - P_j(t) (\lambda_j + \mu_j)] \quad (\text{forward eq.})$$

### LIMIT BEHAVIOR OF CTMC

$$\{X_n : n \in \mathbb{N}_0\} \quad (\text{embedded DTMC}), \quad \underline{\pi} = [\pi_j]_{j \in \mathcal{S}} \quad (\text{stationary distribution of the embedded DTMC})$$

$$P_j = \lim_{t \rightarrow +\infty} P_{ij}(t) \quad P_j = \frac{\pi_j \nu_j}{\sum_{k \in \mathcal{S}} \frac{\pi_k}{\nu_k}}, \quad j \in \mathcal{S}$$

$$\underline{P} = [P_j]_{j \in \mathcal{S}}, \quad \underline{P} \times \mathbf{R} = \underline{0}, \quad \sum_{j \in \mathcal{S}} P_j = 1 \quad P_j \times \nu_j = \sum_{i \in \mathcal{S}} P_i \times q_{ij}, \quad j \in \mathcal{S}$$

$$P_0 = \left[ 1 + \sum_{n=1}^{+\infty} \frac{\lambda_0 \lambda_1 \dots \lambda_{n-1}}{\mu_1 \mu_2 \dots \mu_n} \right]^{-1}, \quad P_j = \frac{\lambda_{j-1}}{\mu_j} P_{j-1} = P_0 \times \frac{\lambda_0 \lambda_1 \dots \lambda_{j-1}}{\mu_1 \mu_2 \dots \mu_j}, \quad j \in \mathbb{N} \quad (\text{birth and death proc.})$$

**Queues**


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$$L_s \quad E(L_s) = \lambda_e E(W_s) = \lambda \times (1 - P_b) \times E(W_s)$$

$$E(L_s) = \frac{\lambda_e}{\mu} + E(L_q)$$

$$L_q \quad E(L_q) = \lambda_e E(W_q)$$

$$W_s \quad E(W_s) = \frac{1}{\mu} + E(W_q)$$


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**M/M/1**


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$$\begin{aligned} \text{Rates} \quad \lambda_k &= \lambda, \quad k \in \mathbb{N}_0 \\ \mu_k &= \mu, \quad k \in \mathbb{N} \quad \left( \rho = \frac{\lambda}{\mu} < 1 \right) \end{aligned}$$


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$$\begin{aligned} L_s \quad P(L_s = k) &= \rho^k (1 - \rho), \quad k \in \mathbb{N}_0 \\ E(L_s) &= \frac{\rho}{(1 - \rho)} \end{aligned}$$


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$$\begin{aligned} L_q \quad P(L_q = k) &= \begin{cases} 1 - \rho^2, & k = 0 \\ \rho^{k+1} (1 - \rho), & k \in \mathbb{N} \end{cases} \\ E(L_q) &= \frac{\rho^2}{(1 - \rho)} \end{aligned}$$


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$$\begin{aligned} W_s \quad (W_s \mid L_s = k) &\sim \text{Gamma}(k + 1, \mu), \quad k \in \mathbb{N}_0 \\ W_s &\sim \text{Exponential}(\mu(1 - \rho)) \\ E(W_s) &= \frac{1}{\mu(1 - \rho)} \end{aligned}$$


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$$\begin{aligned} W_q \quad (W_q \mid L_s = k) &\sim \text{Gamma}(k, \mu), \quad k \in \mathbb{N} \\ &0, \quad t < 0 \\ F_{W_q}(t) &= \begin{cases} 1 - \rho, & t = 0 \\ (1 - \rho) + \rho \times F_{\text{Exp}(\mu(1 - \rho))}(t), & t > 0 \end{cases} \\ (W_q \mid W_q > 0) &\sim \text{Exponential}(\mu(1 - \rho)) \\ E(W_q) &= \frac{\rho}{\mu(1 - \rho)} \end{aligned}$$


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**M/M/ $\infty$** 


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$$\begin{aligned} \text{Rates} \quad \lambda_k &= \lambda, \quad k \in \mathbb{N}_0 \\ \mu_k &= k\mu, \quad k \in \mathbb{N} \quad \left( \rho = \frac{\lambda}{\mu} < +\infty \right) \end{aligned}$$


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$$L_s \quad L_s \sim \text{Poisson}(\lambda/\mu)$$


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$$L_q \stackrel{st}{=} 0, \quad W_s \sim \text{Exp}(\mu), \quad W_q \stackrel{st}{=} 0$$


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$X(t)$  = number of customers in the system at time  $t$

$$(X(t) \mid X(0) = 0) \sim \text{Poisson}(\lambda(1 - e^{-\mu t})/\mu)$$


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**M/G/ $\infty$** 


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$$(X(t) \mid X(0) = 0) \sim \text{Poisson} \left( \lambda \int_0^t [1 - G(t - s)] ds \right)$$

$$\lim_{t \rightarrow +\infty} (X(t) \mid X(0) = 0) \sim \text{Poisson}(\lambda/\mu)$$


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## M/M/m

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Rates	$\lambda_k = \lambda, \quad k \in \mathbb{N}_0$ $\mu_k = \begin{matrix} k\mu, & k = 1, \dots, m \\ m\mu, & k = m + 1, m + 2, \dots \end{matrix} \quad \left(\rho = \frac{\lambda}{m\mu} < 1\right)$
$L_s$	$P(L_s = k) = \begin{matrix} \frac{m!}{k!} (1 - \rho)(m\rho)^{k-m} C(m, m\rho), & k = 0, 1, \dots, m - 1 \\ (1 - \rho) \rho^{k-m} C(m, m\rho), & k = m, m + 1, \dots \end{matrix}$ $C(m, m\rho) = P(L_s \geq m) = \frac{\frac{(m\rho)^m}{m!(1-\rho)}}{\sum_{j=0}^{m-1} \frac{(m\rho)^j}{j!} + \frac{(m\rho)^m}{m!(1-\rho)}}$ $C(1, \rho) = \rho$ $C(2, 2\rho) = \frac{2\rho^2}{1+\rho}$ $E(L_s) = m\rho + \frac{\rho}{1-\rho} C(m, m\rho)$
$L_q$	$P(L_q = k) = \begin{matrix} 1 - \rho C(m, m\rho), & k = 0 \\ (1 - \rho) \rho^k C(m, m\rho), & k \in \mathbb{N} \end{matrix}$ $E(L_q) = \frac{\rho}{1-\rho} C(m, m\rho)$
$W_s$	$(W_s   L_s = k) \sim \begin{matrix} \text{Exp}(\mu), & k = 0, \dots, m - 1, \\ \text{Exp}(\mu) \star \text{Gamma}(k - m + 1, m\mu), & k = m, m + 1, \dots, \end{matrix}$ $1 - F_{W_s}(t) = \begin{cases} [1 + \mu t C(m, m\rho)] e^{-\mu t}, & t \geq 0, \quad \rho = \frac{m-1}{m} \\ \left[1 + \frac{e^{\mu[1-m(1-\rho)]t}}{1-m(1-\rho)} \times C(m, m\rho)\right] e^{-\mu t}, & t \geq 0, \quad \rho \neq \frac{m-1}{m} \end{cases}$ $E(W_s) = \frac{1}{\mu} + \frac{C(m, m\rho)}{m\mu(1-\rho)}$
$W_q$	$(W_q   L_s = k) \sim \text{Gamma}(k - m + 1, m\mu), \quad k = m, m + 1, \dots$ $(W_q   W_q > 0) \sim \text{Exponential}(m\mu(1 - \rho))$ $1 - F_{W_q}(t) = \begin{matrix} 1, & t < 0 \\ C(m, m\rho), & t = 0 \\ C(m, m\rho) \times [1 - F_{\text{Exp}(m\mu(1-\rho))}(t)], & t > 0 \end{matrix}$ $E(W_q) = \frac{C(m, m\rho)}{m\mu(1-\rho)}$

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## M/M/m/m

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Rates	$\lambda_k = \begin{matrix} \lambda, & k = 0, 1, \dots, m - 1 \\ 0, & k = m, m + 1, \dots \end{matrix}$ $\mu_k = \begin{matrix} k\mu, & k = 1, \dots, m \\ 0, & k = m + 1, m + 2, \dots \end{matrix} \quad \left(\rho = \frac{\lambda}{m\mu} < +\infty\right)$
$L_s$	$P(L_s = k) = \begin{matrix} \frac{\frac{(m\rho)^k}{k!}}{\sum_{j=0}^m \frac{(m\rho)^j}{j!}} = \frac{m!}{k! (m\rho)^{m-k}} \times B(m, m\rho), & k = 0, 1, \dots, m \\ 0, & k = m + 1, m + 2, \dots \end{matrix}$ $B(m, m\rho) = P(L_s = m) = \frac{\frac{(m\rho)^m}{m!}}{\sum_{j=0}^m \frac{(m\rho)^j}{j!}}$ $E(L_s) = m\rho [1 - B(m, m\rho)]$
$L_q \stackrel{st}{=} 0, \quad W_s \sim \text{Exp}(\mu), \quad W_q \stackrel{st}{=} 0$	

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