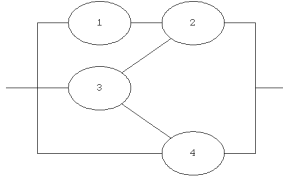


1. Considere que o sistema eléctrico no esquema abaixo é constituído por componentes independentes com fiabilidades iguais a  $p_i$ ,  $i = 1, 2, 3, 4$ .



- (a) Determine a função de estrutura e a fiabilidade do sistema. (3.0)

• Reliabilities of the components

$$p_i, i = 1, \dots, 4$$

• Minimal path sets

$$\mathcal{P}_1 = \{1, 2\}$$

$$\mathcal{P}_2 = \{2, 3\}$$

$$\mathcal{P}_3 = \{4\}$$

$$p^* = 3 \text{ minimal path sets}$$

• Structure function

By considering  $X_i \sim \text{Bernoulli}(p_i)$ ,  $i = 1, \dots, 4$ , and applying result (1.13), we can conclude that the structure function of this coherent system equals:

$$\begin{aligned} \phi(\underline{X}) &= 1 - \prod_{j=1}^{p^*} \left( 1 - \prod_{i \in \mathcal{P}_j} X_i \right) \\ &= 1 - (1 - X_1 X_2)(1 - X_2 X_3)(1 - X_4) \end{aligned}$$

• Reliability

Taking into account that  $X_i \stackrel{\text{indep}}{\sim} \text{Bernoulli}(p_i)$  and  $X_i^k \sim X_i$ ,  $k \in \mathbb{N}$ , we obtain, for  $\underline{p} = (p_1, \dots, p_4)$ :

$$\begin{aligned} r(\underline{p}) &= E[\phi(\underline{X})] \\ &= E[1 - (1 - X_1 X_2)(1 - X_2 X_3)(1 - X_4)] \\ &= 1 - E[(1 - X_1 X_2)(1 - X_2 X_3)] \times E(1 - X_4) \\ &= 1 - E(1 - X_1 X_2 - X_2 X_3 + X_1 X_2^2 X_3) \times E(1 - X_4) \\ &= 1 - E(1 - X_1 X_2 - X_2 X_3 + X_1 X_2 X_3) \times E(1 - X_4) \\ &= 1 - (1 - p_1 p_2 - p_2 p_3 + p_1 p_2 p_3) \times (1 - p_4). \end{aligned}$$

- (b) Calcule a importância da fiabilidade da componente 3 e concretize-a para  $p_i = p$ ,  $i = 1, 2, 3, 4$ . (2.0)

• Importance of the reliability of component 3

$$\begin{aligned} I_r(3) &\stackrel{(1.29)}{=} \frac{\partial r(\underline{p})}{\partial p_3} \\ &= \frac{\partial [1 - (1 - p_1 p_2 - p_2 p_3 + p_1 p_2 p_3) \times (1 - p_4)]}{\partial p_3} \\ &= -(-p_2 + p_1 p_2) \times (1 - p_4) \\ &\stackrel{p_i=p}{=} (p - p^2) \times (1 - p) \\ &= p \times (1 - p)^2. \end{aligned}$$

• Obs.

Let us remind the reader that  $I_r(3) \stackrel{(1.28)}{=} r(1_3, \underline{p}) - r(0_3, \underline{p})$ , where  $(1_3, \underline{p}) = (p_1, p_2, 1, p_4)$ ,  $(0_3, \underline{p}) = (p_1, p_2, 0, p_4)$  and  $r(\underline{p}) = 1 - (1 - p_1 p_2 - p_2 p_3 + p_1 p_2 p_3) \times (1 - p_4)$ . Thus,

$$\begin{aligned} I_r(3) &= [1 - (1 - p_1 p_2 - p_2 + p_1 p_2) \times (1 - p_4)] - [1 - (1 - p_1 p_2) \times (1 - p_4)] \\ &= [(1 - p_1 p_2) - (1 - p_2)] \times (1 - p_4) \\ &= (p_2 - p_1 p_2) \times (1 - p_4) \\ &\stackrel{p_i=p}{=} (p - p^2) \times (1 - p) \\ &= p \times (1 - p)^2. \end{aligned}$$

2. Admita que uma aeronave com 4 motores só será capaz de voar se possuir pelo menos 3 dos motores a funcionar.

- (a) Obtenha um limite inferior e um superior para a fiabilidade da aeronave, assumindo que os 4 motores estão associados (positivamente) e a fiabilidade de cada motor é de 97.5%. (3.0)

• System/components

$$p_i = p = 0.975, i = 1, \dots, 4$$

Since the 3-out-of-4 system is coherent and its components operate in a positively associated fashion, we can apply Theorem 1.70, namely result (1.42), after having identified the minimal path sets and minimal cut sets of this system.

• Minimal path sets

$$\mathcal{P}_1 = \{1, 2, 3\}$$

$$\mathcal{P}_2 = \{1, 2, 4\}$$

$$\mathcal{P}_3 = \{1, 3, 4\}$$

$$\mathcal{P}_4 = \{2, 3, 4\}$$

$$p^* = 4 \text{ minimal path sets}$$

• Minimal cut sets

$$\mathcal{K}_1 = \{1, 2\}$$

$$\mathcal{K}_2 = \{1, 3\}$$

$$\mathcal{K}_3 = \{1, 4\}$$

$$\mathcal{K}_4 = \{2, 3\}$$

$$\mathcal{K}_5 = \{2, 4\}$$

$$\mathcal{K}_6 = \{3, 4\}$$

$q = 6$  minimal cut sets

• **Lower bound for the reliability**  $r(\underline{p})$

$$\begin{aligned} r(\underline{p}) &\stackrel{(1.42)}{\geq} \max_{j=1, \dots, p^*} \prod_{i \in \mathcal{P}_j} p_i \\ &\stackrel{p_i = p}{=} \max_{j=1, \dots, p^*} p^{\#\mathcal{P}_j} \\ &\stackrel{\#\mathcal{P}_j = 3, \forall j}{=} p^3 \\ &\stackrel{p = 0.975}{=} 0.975^3 \\ &\simeq 0.926859. \end{aligned}$$

• **Upper bound for the reliability**

$$\begin{aligned} r(\underline{p}) &\stackrel{(1.42)}{\leq} \min_{j=1, \dots, q} \left[ 1 - \prod_{i \in \mathcal{K}_j} (1 - p_i) \right] \\ &\stackrel{p_i = p}{=} \min_{j=1, \dots, q} \left[ 1 - (1 - p)^{\#\mathcal{K}_j} \right] \\ &\stackrel{\#\mathcal{K}_j = 2, \forall j}{=} 1 - (1 - p)^2 \\ &\stackrel{p = 0.975}{=} 1 - (1 - 0.975)^2 \\ &= 0.999375. \end{aligned}$$

• **Obs.**

Since the components are positively associated we cannot apply Theorem 1.68 (for independent components!) which sometimes provides stricter bounds than the ones we just obtained by using Theorem 1.70.

(b) Admita agora que os tempos até falha dos 4 motores (em milhares de horas) são *i.i.d.* com distribuição Exponencial e parâmetro de escala  $\lambda^{-1} = 2$  e obtenha a função de fiabilidade da aeronave para um período de 1000 h. (2.0)

• **Individual durations of the components of the 3 – out – of – 4 system**

$$\begin{aligned} T_i &\stackrel{i.i.d.}{\sim} \text{Exponential}(\lambda = 0.5), \quad i = 1, \dots, 4 \\ F_i(t) &= F(t) = 1 - e^{-0.5t}, \quad t \geq 0 \end{aligned}$$

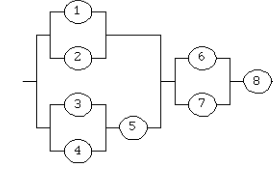
• **Duration of the 3 – out – of – 4 system**

$$\begin{aligned} T &= \text{duration of the 3 – out – of – 4 system} \\ T &\stackrel{(2.6)}{=} T_{(n-k+1)}^{n=4, k=3} T_{(2)} \end{aligned}$$

• **Reliability function of  $T$**

$$\begin{aligned} R_T(t) &\stackrel{(2.8)}{=} F_{\text{Binomial}(n, F(t))(n-k)} \\ &= F_{\text{Binomial}(4, 1 - e^{-0.5t})(4-3)} \\ &= \sum_{m=0}^1 \binom{4}{m} (1 - e^{-0.5t})^m [1 - (1 - e^{-0.5t})]^{4-m} \\ &= e^{-2t} + 4(1 - e^{-0.5t})e^{-1.5t} \\ &\stackrel{t=1}{\simeq} 0.486515 \end{aligned}$$

3. O esquema seguinte reporta-se a uma estrutura constituída por componentes com durações *i.i.d.* e *IHR*.



(a) O que pode concluir acerca do comportamento monótono da sua função taxa de falha da duração de vida da estrutura? Justifique. (2.0)

• **Individual durations**

$$T_i \stackrel{i.i.d.}{\sim} \text{IHR}, \quad i = 1, \dots, 8$$

• **Duration of the system**

Let:

$$\begin{aligned} Z_1 &= \max\{\max\{T_1, T_2\}, \min\{\max\{T_3, T_4\}, T_5\}\}; \\ Z_2 &= \max\{T_6, T_7\}; \\ Z_3 &= T_8. \end{aligned}$$

Then

$$T = \min\{Z_1, Z_2, Z_3\}.$$

• **Stochastic ageing of  $T$**

Since  $T_i \stackrel{i.i.d.}{\sim} \text{IHR}$ ,  $i = 1, \dots, 8$ , we can add that  $Z_3 = T_8 \in \text{IHR}$ , and we can apply Proposition 3.23 and successively get:

- $\max\{T_1, T_2\}, \max\{T_3, T_4\} \in \text{IHR}$  by result (3.14);
- $\min\{\max\{T_3, T_4\}, T_5\} \in \text{IHR}$  by result (3.11);
- $Z_2 = \max\{T_6, T_7\} \in \text{IHR}$  by result (3.14);

However, according to result (3.13), the maximum of two independent (but not identically distributed) *IHR* r.v. is not necessarily an *IHR* r.v., therefore we cannot infer the stochastic ageing of  $Z_1$  or the one of  $T = \min\{Z_1, Z_2, Z_3\}$ .

(b) Estabeleça um limite inferior para a função de fiabilidade da sub-estrutura constituída pelas componentes 6 e 7. (3.0)

• **Common reliability function of components 6 and 7**

$$R(t) = P(T_6 > t) = P(T_7 > t)$$

• **Lower bound for  $R(t)$**

Recall that

- $T_i \stackrel{i.i.d.}{\sim} \text{IHR}$ ,  $i = 6, 7$ ,

and admit that

- $E(T_i) = \mu$ ,  $i = 6, 7$ .

Then we can apply Corollary 3.47 and state that

$$R(t) \geq \begin{cases} e^{-\frac{t}{\mu}}, & t \leq \mu \\ 0, & t > \mu. \end{cases}$$

- **Reliability function of components**  $Z_2 = \max\{T_6, T_7\}$

$$R_{Z_2}(t) = P[\max\{T_6, T_7\} > t] \\ \stackrel{(2.5)}{=} 1 - [1 - R(t)]^2$$

- **Lower bound for  $R_{Z_2}(t)$**

Capitalizing on the lower bound for  $R(t)$ , we can add that

$$[1 - R(t)]^2 \leq \begin{cases} \left(1 - e^{-\frac{t}{\mu}}\right)^2, & t \leq \mu \\ (1 - 0)^2 = 1, & t > \mu; \end{cases} \\ 1 - [1 - R(t)]^2 = R_{Z_2}(t) \\ \geq \begin{cases} 1 - \left(1 - e^{-\frac{t}{\mu}}\right)^2 = e^{-\frac{t}{\mu}} \times \left(2 - e^{-\frac{t}{\mu}}\right), & t \leq \mu \\ 1 - 1 = 0, & t > \mu. \end{cases}$$

4. O registro dos instantes de ocorrência de falhas (em horas) de 10 impressoras colocadas simultaneamente em teste conduziu às seguintes observações: 43, 55, 84, 172, 215, 419, 565, 586, 687, 711.

- (a) Calcule uma estimativa da função taxa de falha da duração de uma impressora para  $t = 55$  h. (1.5)

- **Ordered data**

(43, 55, 84, 172, 215, 419, 565, 586, 687, 711)

- **Nonparametric estimate of  $\lambda(t_{(i)})$**

We are dealing with complete ungrouped failure data, thus, we can use the nonparametric estimate of  $\lambda(t_{(i)})$  in Table 5.1,

$$\lambda(t_{(i)}) = \frac{1}{n - i + 0.625} \times \frac{1}{t_{(i+1)} - t_{(i)}},$$

and get

$$\lambda(55) = \lambda(t_{(2)}) \\ = \frac{1}{10 - 2 + 0.625} \times \frac{1}{84 - 55} \\ \simeq 0.003998.$$

Admita agora que as 10 observações dizem respeito a 10 impressoras de entre um grupo de 30 colocadas simultaneamente em teste e que as durações das impressoras são v.a. i.i.d. com distribuição Exponencial.

- (b) Reavalie a estimativa que obteve em (a) e determine um intervalo de confiança a 95% para a fiabilidade de uma impressora para um período de 200 h. (3.5)

- **Distribution assumption**

$$T_i \stackrel{i.i.d.}{\sim} T \sim \text{Exponential}(\lambda), i = 1, \dots, n.$$

- **Unknown parameter**

$$\lambda_T(t) = \lambda, t > 0.$$

- **Life test**

Since the end of the test was determined by the  $r = 10^{\text{th}}$  failure and nothing in this exercise suggests that the  $n = 20$  rubber gaskets were replaced during the life test, we are dealing with a

- Type II/item censored testing without replacement.

- **Censored data**

$n = 30$

$r = 10$

$(t_{(1)}, \dots, t_{(r)}) = (43, 55, 84, 172, 215, 419, 565, 586, 687, 711)$

- **Cumulative total time in test**

According to Definition 5.17, the cumulative total time in test is given by:

$$\tilde{t} = \sum_{i=1}^r t_{(i)} + (n - r) \times t_{(r)} \\ = (43 + \dots + 711) + (30 - 10) \times 711 \\ = 17757.$$

- **Estimates of  $\lambda$**

According to Table 5.11, the ML of  $\lambda$  is equal to

$$\hat{\lambda} = \frac{r}{\tilde{t}} \\ = \frac{10}{17757} \\ \simeq 0.000563158.$$

- **Another unknown parameter**

$R_T(t) = e^{-\lambda t}$ ,  $t = 200$ .

- **Confidence interval for  $\lambda$**

According to Table 5.16 of the lecture notes,

$$CI_{(1-\alpha) \times 100\%}(\lambda) = [\lambda_L; \lambda_U] \\ = \left[ \frac{F_{\chi_{(2r)}^2}^{-1}(\alpha/2)}{2 \times \tilde{t}}; \frac{F_{\chi_{(2r)}^2}^{-1}(1 - \alpha/2)}{2 \times \tilde{t}} \right] \\ CI_{95\%}(\lambda) = \left[ \frac{F_{\chi_{(20)}^2}^{-1}(0.025)}{2 \times 17757}; \frac{F_{\chi_{(20)}^2}^{-1}(0.975)}{2 \times 17757} \right] \\ = \left[ \frac{9.591}{2 \times 17757}; \frac{34.17}{2 \times 17757} \right] \\ \simeq [0.000270063; 0.000962156].$$

- **Confidence interval for  $R_T(t)$**

Since  $R_T(t) = e^{-\lambda t}$  is a decreasing function of  $\lambda > 0$ , we can state that

$$CI_{95\%}(e^{-\lambda t}) = [e^{-\lambda_U t}; e^{-\lambda_L t}] \\ \stackrel{t=200}{\simeq} [0.824951; 0.947420].$$